

Goal: Smoother Density Estimation.

3 Kernel Density Estimation

Recall the definition of a density function

$$f(y) \equiv \frac{d}{dy} F(y) \equiv \lim_{h \rightarrow 0} \frac{F(y+h) - F(y-h)}{2h} = \lim_{h \rightarrow 0} \frac{F(y+h) - F(y)}{h},$$

where $F(x)$ is the cdf of the random variable Y .

small
take this and approximate w/ fixed h ,
use Ecdf.

$$\hat{f}(x) = \frac{\hat{F}_n(x+h) - \hat{F}_n(x)}{h} \quad \text{histogram.}$$

What if instead, we replace $F(y+h) - F(y-h)$ with ecdf

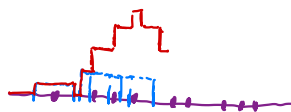
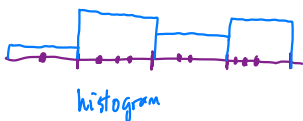
$$\begin{aligned} \Rightarrow \hat{f}(y) &= \frac{\hat{F}_n(y+h) - \hat{F}_n(y-h)}{2h} = \frac{\#\{y_i \in (y-h, y+h]\}}{2nh} \\ &= \frac{\sum_{i=1}^n \mathbb{I}(y_i \in (y-h, y+h])}{2nh} \end{aligned}$$

$$= \frac{1}{nh} \sum_{i=1}^n \frac{1}{2} \mathbb{I}(y-h < y_i \leq y+h)$$

$$= \frac{1}{nh} \sum_{i=1}^n \frac{1}{2} \mathbb{I}\left(-1 < \frac{y-y_i}{h} \leq 1\right)$$

$$= \frac{1}{nh} \sum_{i=1}^n K\left(\frac{y-y_i}{h}\right) \quad \text{where } K \text{ is a Uniform density on } [-1, 1]$$

kernel function.



↑
still not continuous (because Uniform density not continuous!)

⇒ another kernel may lead to smoother estimate.

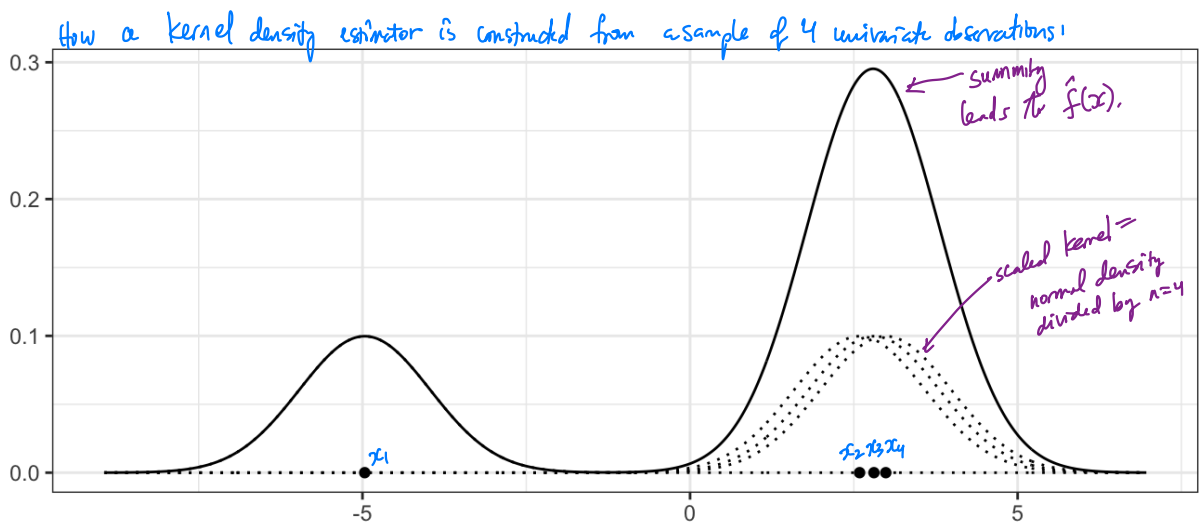
A kernel function assigns weights to the contribution given by each y_i to $\hat{f}(y)$, depending on proximity to y .

This will weight all points within h of y equally. A univariate kernel density estimator will allow a more flexible weighting scheme.

Typically, kernel functions are positive everywhere and symmetric about zero.

Examples of ideas for such functions? Normal density, Student t (others exist).

Additionally, constraining K so that $\int z^2 K(z) dz = 1$ allows h to play role of scale parameter (not required).

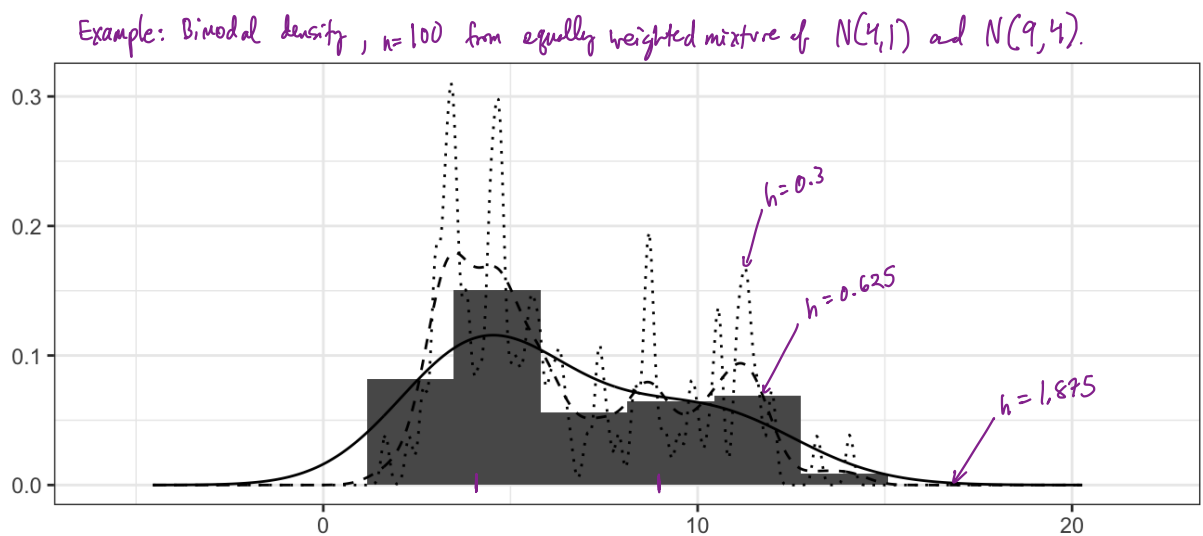


3.1 Choice of Bandwidth

The bandwidth parameter controls the smoothness of the density estimate. *for a given kernel.*

bandwidth determines tradeoff bet/ bias and variance.

The tradeoff that results from choosing the bandwidth + kernel can be quantified through a measure of accuracy of \hat{f} , such as MISE.



For large h , oversmoothing (lose 2nd mode).

For small h , undersmoothing (many false modes).

To understand bandwidth selection, let us analyze ^(AMISE) MISE. Suppose that K is a symmetric, continuous probability density function with mean 0 and variance $0 < \sigma_K^2 < \infty$. Let $R(g) = \int g^2(z)dz$. Recall that

$$\text{MISE} = \int \text{MSE}(\hat{f}(x))dx =$$

Now let $h \rightarrow 0$ and $nh \rightarrow \infty$ as $n \rightarrow \infty$.

To minimize AMISE with respect to h ,

The term $R(f'')$ measures the roughness of the true underlying density. In general, rougher densities are more difficult to estimate and require smaller bandwidth.

The term $[\sigma_K R(K)]^{4/5}$ is a function of the kernel function K .

3.1.1 Cross Validation

3.1.2 Plug-in Methods

If the reference density f is Gaussian and a Gaussian kernel K is used,

Empirical estimation of $R(f'')$ may be a better option.

3.2 Choice of Kernel

There are two choices we have to make to perform density estimation:

3.2.1 Epanechnikov Kernel

The *Epanechnikov kernel* results from choosing K to minimize $[\sigma_K R(K)]^{4/5}$, restricted to be a symmetric density with finite moments and variance equal to 1

3.2.2 Canonical Kernels

Unfortunately a particular value of h corresponds to a different amount of smoothing depending on which kernel is being used.

Let h_K and h_L denote the bandwidths that minimize AMISE when using symmetric kernel densities K and L . Then,

Suppose we rescale a kernel shape so that $h = 1$ corresponds to a bandwidth of $\delta(K)$,

3.3 Bootstrapping and Variability Plot

